



$$\boxed{1} \quad \boxed{a} \quad E(X^{1/2}) \leq (E(X))^{1/2} \quad \text{for } X > 0.$$

$$\begin{aligned} g(x) &= x^{1/2} \\ \Rightarrow g'(x) &= \frac{1}{2} x^{-1/2} \\ \Rightarrow g''(x) &= -\frac{1}{4} x^{-3/2} < 0 \end{aligned} \quad \begin{aligned} &\Rightarrow g(x) \text{ is concave} \\ &\Rightarrow E(g(x)) \leq g(E(x)) \quad \text{by Jensen's Inequality} \end{aligned}$$

$$\boxed{b} \quad X \sim \text{Beta}(\alpha, \beta) \Rightarrow$$

$$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \quad \text{for } \alpha, \beta > 0, \quad 0 < x < 1$$

$$= \underbrace{\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}}_{c(\theta)} \underbrace{\frac{1}{[0,1]}}_{h(x)} \exp \left[\underbrace{(\alpha-1)}_{w_1(\theta)} \log(x) + \underbrace{(\beta-1)}_{w_2(\theta)} \log(1-x) \right]$$

$\uparrow \quad \uparrow \quad \uparrow \quad \uparrow$
 $w_1(\theta) \quad t_1(x) \quad w_2(\theta) \quad t_2(x)$

Yes.

$$\boxed{c} \quad X \sim \text{Exp}(\theta) \Rightarrow f(x) = \frac{1}{\theta} e^{-x/\theta} \quad \text{for } x > 0$$

Location shifted:

$$f(x-\mu) = \frac{1}{\theta} e^{-\frac{(x-\mu)}{\theta}} \quad \text{for } x-\mu > 0$$

$$= \boxed{\frac{1}{\theta} e^{-\frac{(x-\mu)}{\theta}} \quad \text{for } x > \mu}$$

$$\boxed{d} \quad E(Y) = E(E(Y|X))$$

$$\begin{aligned} E(E(Y|X)) &= \int \left(\int y f(y|x) dy \right) f(x) dx \\ &= \iint y f(x,y) dy dx \\ &= \int y \int f(x,y) dx dy \\ &= \int y f(y) dy \\ &= E(Y) \quad \square \end{aligned}$$

$$\boxed{e} \quad P(e^{tX} \geq 1) \leq M_X(t) = E(e^{tX})$$

From Chebyshev's inequality:

$$P(g(X) \geq r) \leq \frac{E(g(X))}{r} \quad \text{if } g(x) > 0, r > 0$$

In the above, $g(x) = e^{tX} > 0$

$$r = 1.$$

□

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$$X + Y, \quad X, Y \sim \text{Gamma}(\alpha, \beta)$$

$$u = \frac{X}{X+Y} \quad v = X+Y$$

a) $f(u, v) = ?$

$$f(x, y) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} \cdot \frac{1}{\Gamma(\alpha)\beta^\alpha} y^{\alpha-1} e^{-y/\beta} \quad \text{for } x, y > 0$$

$$u = \frac{x}{x+y} \Rightarrow u = \frac{v-y}{v-y+y} = \frac{v-y}{v} \Rightarrow u = 1 - \frac{y}{v} \Rightarrow \frac{y}{v} = 1-u \Rightarrow y = v(1-u)$$

$$v = x+y \Rightarrow x = v-y \Rightarrow x = \cancel{v-y} \cdot v - v(1-u) = v - v + uv = uv \Rightarrow x = uv$$

$$\begin{aligned} \frac{dx}{du} &= v & \frac{dx}{dv} &= u \\ \frac{dy}{du} &= -v & \frac{dy}{dv} &= 1-u \end{aligned} \Rightarrow |J| = |v(1-u) + vu| = |v - uv + uv| = v$$

$$\Rightarrow f(u, v) = \frac{1}{\Gamma(\alpha)\beta^\alpha} (uv)^{\alpha-1} e^{-uv/\beta} \frac{1}{\Gamma(\alpha)\beta^\alpha} (v(1-u))^{\alpha-1} e^{-v(1-u)/\beta} v$$

$$= \frac{1}{\Gamma(\alpha)^2 \beta^{2\alpha}} u^{\alpha-1} (1-u)^{\alpha-1} v^{2(\alpha-1)+1} \exp\left(-\frac{1}{\beta}(uv + v(1-u))\right)$$

$$= \frac{1}{\Gamma(\alpha)^2 \beta^{2\alpha}} u^{\alpha-1} (1-u)^{\alpha-1} v^{2\alpha-1} e^{-\frac{v}{\beta}} \quad \text{for } v > 0, 0 < u < 1$$

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$$\begin{aligned}
 \text{b) } f(v) &= \int_0^1 f(u, v) du \\
 &= \frac{1}{\Gamma(\alpha)^2 \beta^{2\alpha}} v^{2\alpha-1} e^{-v/\beta} \frac{\Gamma(\alpha)\Gamma(\alpha)}{\Gamma(\alpha+\alpha)} \int_0^1 \frac{\Gamma(\alpha+\alpha)}{\Gamma(\alpha)\Gamma(\alpha)} u^{\alpha-1} (1-u)^{\alpha-1} du \\
 &= \frac{\Gamma(\alpha)^2}{\Gamma(\alpha)^2 \Gamma(2\alpha) \beta^{2\alpha}} v^{2\alpha-1} e^{-v/\beta} \\
 &= \frac{1}{\Gamma(2\alpha) \beta^{2\alpha}} v^{2\alpha-1} e^{-v/\beta} \quad \text{for } v > 0 \\
 &\Rightarrow V \sim \text{Gam}(2\alpha, \beta)
 \end{aligned}$$

$$\begin{aligned}
 \text{c) } f(u) &= \int_0^{\infty} f(u, v) dv \\
 &= \frac{1}{\Gamma(\alpha)^2 \beta^{2\alpha}} u^{\alpha-1} (1-u)^{\alpha-1} \frac{\Gamma(2\alpha) \beta^{2\alpha}}{1} \int_0^{\infty} \frac{1}{\Gamma(2\alpha) \beta^{2\alpha}} v^{2\alpha-1} e^{-v/\beta} dv \\
 &= \frac{\Gamma(2\alpha)}{\Gamma(\alpha)\Gamma(\alpha)} u^{\alpha-1} (1-u)^{\alpha-1} \quad \text{for } 0 < u < 1 \\
 &\Rightarrow U \sim \text{Beta}(\alpha, \alpha)
 \end{aligned}$$

d) Yes, because $f(u, v) = f(u)f(v)$,
 or more simply $f(u, v) = h(u)g(v)$.
 - The joint distribution factors.