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Bios 312: Modern Regression Analysis

Midterm Examination

February 28, 2013

Name:KEY		
difficulty, so you		answers to all questions. Questions are of varying levels of geous to skip questions you find especially difficult, and return xam.
You are allowed	one (1) page of your	own notes to assist you when taking the exam.
use at least four s	significant digits; rep	th arithmetic. When making intermediate calculations, always port at least three significant digits. If you are running short of orm to receive the majority of credit.
		lieve cannot be answered without making additional able assumption that you make, and proceed.
		e. If you are unable to truthfully sign the pledge for any reason, cuss the circumstances with the instructor.
PLEDGE: On my	honor, I have neithe	r given nor received unauthorized aid on this examination
Signature:		
This exam contotal points.	sists of 11 pages	including an Appendix of Results. There are 140
· ·	on 1: Parts 1-7, 60	•
•	on 2: Parts 1-4, 40	•
• Question	on 3: Parts 1-4, 40	0 points
Stem and Lea	of Plot of Grade I	Distribution
	0 4 1 2	Mean: 105 Median: 106
	35667	Median. 100

Question 1 (60 points) Use the results in the Appendix to answer the following questions about the association of serum arginine (arg) with cholesterol (chol), age, and male sex. Assume the assumption of homoscedasticity holds for all models.

1-1. Based on the simple linear regression model (only including cholesterol as covariate), what scientific conclusions can you make about the association between cholesterol and arginine? Provide an interpretation for the cholesterol coefficient. Include estimates of the effect size and statistical inference in your written description of the results. (10 points)

Comparing two subjects who differ in their cholesterol level by one unit, a 1-unit increase in cholesterol is associated with a 3.46 unit change in serum arginine. We are 95% confident that the true change in arginine per unit change in cholesterol is between 1.22 and 5.68. Because the confidence interval does not contain 0 (and p = 0.003), there is a significant association between cholesterol and arginine at the 5% significance level.

1-2. Based on the multivariable linear regression model with cholesterol, age and male as covariates, what scientific conclusions can you make about the association between cholesterol and arginine? Provide an interpretation for the cholesterol coefficient. Include estimates of the effect size and statistical inference in your written description of the results. (10 points)

Comparing two subjects who differ in their cholesterol level by one unit but have the same age and gender, a 1-unit increase in cholesterol is associated with a 2.35 unit increase in serum arginine. We are 95% confident that the true change in arginine per unit change in cholesterol is between 0.36 and 4.35. Because the confidence interval does not contain 0 (and p = 0.02), there is a significant association between cholesterol and arginine at the 5% significance level.

1-3. Is there any evidence that age confounds the association between arginine and cholesterol? Do you consider age to be a precision variable? Explain your answers. (10 points)

A confounder must be related to both the outcome (causally) and the predictor of interest. Symptoms of confounding in a linear regression model are scientifically meaningful (e.g. 10%) changes in the regression coefficient estimates. When we compare the coefficient for cholesterol in the unadjusted model to the model where we adjust for age, the coefficient changes from 3.46 to 2.40; this large change is indicative of confounding. To evaluate a precision variable, we should look at the standard error of the cholesterol coefficient. It changes from 1.12 (unadjusted) to 1.28 (adjusted); because it increases, there is little evidence that it is a precision variable.

1-4. Is there any evidence that male sex confounds the association between arginine and cholesterol? Do you consider male sex to be a precision variable? Explain your answers. (10 points)

Using the same criteria as above, male sex appears to be a precision variable. When we adjust for male sex, the standard error of the cholesterol coefficient decreases from 1.12 (unadjusted) to 0.89 (adjusted for male); also note the decrease in the RMSE from 14.5 to 11.5. The estimate of the regression coefficient changes very little (3.46 vs 3.44), so there are not strong symptoms of confounding. Note that this is not definitive proof of lack of confounding as confounding is primarily a concept that does not have a formal statistical test.

1-5. Based on the multivariable linear regression model with cholesterol, age and male as covariates, what is the expected value of arginine when cholesterol is 185, sex is male, and age is 50? What assumption about the form of the relationship between arginine and cholesterol is needed for this to be interpreted as either an expected value or a predicted value of a new observation? (10 points)

-422.97 + 185*2.356 + 50*0.3883 + 17.82 = 50.1

We assume that the relationship between arginine and cholesterol follows a straight line.

1-6. Briefly describe what the assumption of homoscedasticity means for the multivariable linear regression model with cholesterol, age and male as covariates. (5 points)

Homoscedasticity means that the residuals have constant variance for all expected values of Y. Since the expected value is a function of age, cholesterol and male, the variance will be the same across groups defined by these three covariates.

1-7. Suppose you fit the simple linear regression model (only age as a covariate) and use robust standard error estimates. How would this alter your model results? In particular, discuss changes to your estimates of the slope and the standard error of the slope. You are not given this output, so just discuss the changes you expect to occur. (5 points)

The regression coefficient estimates would not change. The standard error estimates would increase because we are told to assume that the assumption of homoscedasticity holds. When this assumption holds, the classical model will be more efficient than using robust standard errors.

Question 2 (40 points): Suppose that we are interested in examining the association of time to death with age and smoking status in a group of elderly patients enrolled in a cohort study. Use the analysis conducted in the appendix to answer the following questions.

2-1) Based on the logistic regression model, what scientific conclusions can you make about the association between age and death? If appropriate, include estimates of the effect size and statistical inference in your written description of the results. (10 points)

No scientifically relevant conclusions can be drawn. This is censored survival data, so logistic regression is not appropriate because it does not take into account the variable time under observation.

(Aside: In this dataset-- although you would have no way of knowing this-- subjects with the earliest censoring times were minorities because a second cohort of individuals was recruited to enhance minority representation. In this case, failure to adjust for the censoring could lead to very different estimates of survival probabilities if age and/or its association with survival differed across race/ethnicity. There are many nuances here, but the general point remains: with censored observations, you need to consider censoring to draw valid scientific conclusions).

2-2) Based on the simple PH regression model (only age as a covariate), what scientific conclusions can you make about the association between age and time to death? If appropriate, include estimates of the effect size and statistical inference in your written description of the results. (10 points)

Every one unit increase in age is associated wit a 1.03 fold increase in the instantaneous risk of death. We are 95% confident that the true hazard ratio is between 1.01 and 1.06. Since the confidence interval does not contain 1, and p = 0.014, we can conclude that there is a significant association between age and the instantaneous risk of death.

2-3) Based on the multivariable PH regression model (age and smoker as covariates), what scientific conclusions can you make about the association between age and time to death? If appropriate, include estimates of the effect size and statistical inference in your written description of the results. (10 points)

Among subjects with the same smoking status but differing in age, the subject who is 1 year older is at a .04 fold increased risk (hazard) of death compared to the younger subject. We are 95% confident that the true adjusted hazard ratio is between 1.01 and 1.07. Since the confidence interval does not contain 1 (and p = 0.008), we conclude that there is a significant association between age and the risk of death, holding smoking status constant.

2-4) Compare the simple and multivariable PH regression models. Is there evidence that smoking status confounds the association between age and time to death? Is there evidence that it is a precision variable? Justify your answer. (10 points)

To evaluate symptoms of confounding, we evaluate if the hazard ratio estimate changes by a scientifically important amount from the unadjusted to the adjusted model. Here, it changes from 1.034 to 1.037, so there is little evidence of confounding. When modeling hazards (or probabilities), precision variables will tend to drive the coefficient estimate away from the null. In this case, we see the hazard ratio increase so there is some evidence that smoking status is a precision variable. Furthermore, smoking status is significantly associated with the risk of death (p = 0.020), which is also consistent with smoking being a precision variable.

Note that I am primarily concerned with your justification of your answer, which should raise the issues stated above. If, for example, you followed the same logic and concluded that smoker was not a precision variable, I would consider that acceptable in this case. In real observational data, it is rare to find that a variable acts purely as either a precision variable or a confounder.

	Females		Mal	es	All Subjects		
	No Disease	Disease	No Disease	Disease	No Disease	Disease	
Not Diabetic	70	24	52	19	122	43	
Diabetic	12	8	12	11	24	19	
Total	82	32	64	30	146	62	

Question 3 (40 points): The following table presents the cross-classification table of sex, diabetic status, and presence of atherosclerotic disease.

We will consider three models for p, the probability of atherosclerotic disease

- Model 3.A: $\log(p/(1-p)) = \alpha_0 + \alpha_1 * male$
- Model 3.B: $\log(p/(1-p)) = \beta_0 + \beta_1 * \text{male} + \beta_2 * \text{diabetic}$
- Model 3.C: $\log(p/(1-p)) = \gamma_0 + \gamma_1 * \text{male} + \gamma_2 * \text{diabetic} + \gamma_3 * \text{diabetic} * \text{male}$
- 3-1) Before considering the models, calculate the following conditional probabilities using the data in the table. Report as either fractions or decimals. Note that the '|' symbol should be read as 'given'; e.g. Prob(Disease | Male) is "The probability of disease given male". (10 points)

Prob(Disease | Male) =
$$30/(30+64) = 0.3191$$

Prob(Disease | Female) =
$$32/(32+82) = 0.2807$$

Prob(Disease | Male and Diabetic) =
$$11/(11+12) = 0.4782$$

Prob(Disease | Female and Diabetic) =
$$8/(8+12) = 0.4000$$

3-2) Suppose we want to fit a simple logistic regression model (Model 3.A) with atherosclerotic disease (0=No Disease, 1=Disease) as the outcome and an indicator for male sex (0=Female, 1=Male) as the predictor. Using the above table, can you find the intercept and slope for such a logistic regression model? If so, what are the estimates of the intercept and slope? If not, explain the difficulty. (10 points)

 α_{θ} is the log odds of disease in females. Using the results from part 1, the odds of disease in females is 0.2807/(1-0.2807) = 0.3902; the log odds is -0.941. This is α_{θ}

 $\alpha_0 + \alpha_1$ is the log odds of disease in males. Using the results from part 1, odds in males =0.3191 / (1-0.3191) = 0.4686; the log odds is -0.7579. This is $\alpha_0 + \alpha_1$, α_1 is thus 0.183.

Another solution, among many, would be to realize that $\exp(\alpha_1)$ is the odds ratio comparing males to female, so α_1 is the log odds ratio. $\log(82*30/(32*64)) = 0.183$

3-3) Suppose we want to fit a multiple logistic regression model (Model 3.B) with atherosclerotic disease as the outcome, and indicators for male sex and diabetic status (0=Not Diabetic; 1 = Diabetic) as the adjustment variables. Using the above table, can you find the intercept and slopes for such a logistic regression model? If so, what are the estimates of the intercept and slopes? If not, explain the difficulty. (10 points)

These estimates cannot be obtained from the given data. In this model β_1 is the change in the log odds of disease comparing a male to a female subject while holding diabetic status constant. Thus, β_1 is a approximately a weighted average of male effect in non-diabetics and the male effect in diabetics. Without knowing the precise weight used by logistic regression, the estimate β_1 of cannot be computed. Similarly β_2 is the change in the log odds comparing diabetics to non-diabetics, averaging over smoking status; without knowing the weights, we cannot compute β_2 .

3-4) Suppose we want to fit a multiple logistic regression model (Model 3.C) with atherosclerotic disease as the outcome, and indicators for male sex, diabetic status, and the interaction of male sex and diabetic status as the adjustment variables. Using the above table, can you find the intercept and slopes for such a logistic regression model? If so, what are the estimates of the intercept and slopes? If not, explain the difficulty. (10 points)

Yes, they can be determined. Parameters in the model have the following interpretations:

```
\gamma_0 is the log odds of disease in female non-diabetics

= log(0.2553/1-.2553) = -1.071
\gamma_0 + \gamma_1 is the log odds of disease in male non-diabetics

= log(0.2676/1-.2676) = -1.007
\gamma_0 + \gamma_2 is the log odds of disease in female diabetics

= log(0.4/1-.4) = -0.4054
\gamma_0 + \gamma_1 + \gamma_2 + \gamma_3 is the log odds of disease in male diabetics

= log(0.4782/1-.4782) = -0.0872
```

The above calculations use the conditional probabilities in part 1, plus two additional probabilities calculated in a similar manner. Using these results, we find that

$$\gamma_0 = -1.07$$
 $\gamma_1 = -1.007 - (-1.071) = 0.064$
 $\gamma_2 = -0.4054 - (-1.071) = 0.665$
 $\gamma_3 = -0.0872 - (-1.071 + 0.064 + 0.6656) = 0.255$

Appendix for Question 1

. tabstat chol age male, stat(n mean sd min p25 p50 p75 max) format(\$9.2f)

stats		chol	age	male
N mean sd min p25 p50 p75 max	-+-	75.00 185.39 1.51 182.03 184.11 185.48 186.46 189.17	75.00 53.41 8.30 40.28 46.97 54.55 60.31 69.76	75.00 0.45 0.50 0.00 0.00 1.00

. regress arg chol

Source	SS	df	MS		Number of obs F(1, 73)		75 9.55
Model Residual	15383.8822	73	2011.9987 210.73811	2	F(1, 73) Prob > F R-squared Adj R-squared	= =	0.0028 0.1157
Total	17395.8809				Root MSE		14.517
arg			 Err.	t P> t	[95% Conf.	In	terval]
chol _cons	3.455854	1.118	441 3.	09 0.003	1.226804 -1011.293	-	.684904 84.7571

. regress arg chol age

Source	SS	df	MS		Number of obs	
Model Residual + Total	2563.65671 14832.2242 17395.8809	72 206.	.82835 003114 079472		F(2, 72) Prob > F R-squared Adj R-squared Root MSE	= 0.0032 = 0.1474
arg	Coef.	Std. Err.	t	P> t	[95% Conf.	<pre>Interval]</pre>
chol age _cons	2.396671 .3809583 -422.0045	1.281303 .2327981 231.5209	1.87 1.64 -1.82	0.065 0.106 0.072	1575602 0831162 -883.533	4.950903 .8450328 39.52406

Appendix for Question 1 (cont.)

. regress arg chol age male

Source	SS	df	MS		Number of obs F(3, 71)	
Model Residual	8465.65039 8930.23055		1.88346		Prob > F R-squared Adj R-squared	= 0.0000 = 0.4866
Total	17395.8809	74 235	.079472		Root MSE	= 11.215
arg	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
chol age male _cons	2.356197 .3882987 17.82004 -422.9713	1.001209 .1819082 2.601426 180.9071	2.35 2.13 6.85 -2.34	0.021 0.036 0.000 0.022	.3598431 .0255841 12.63294 -783.6898	4.352552 .7510133 23.00713 -62.25271

. regress arg chol male

Source	SS	df	MS		Number of obs F(2, 72)	
Model Residual + Total			.990726		Prob > F R-squared Adj R-squared Root MSE	= 0.0000 = 0.4537
arg		Std. Err.			[95% Conf.	Interval]
chol male _cons	3.435826	.8851478 2.664854 164.1075	3.88 6.67 -3.67	0.000 0.000 0.000	1.671316 12.47504 -929.5176	5.200336 23.09962 -275.2329

Appendix for Question 2

The analysis of time to death with age and smoking status consider the following variables:

- *death*: Indicator of death status (1=died on study; 0=alive at end of study)
- *ttodth*: Number of days from enrollment to either death or the end of the study
- age: Age in years
- *smoker*: Current smoking status (1=Current smoker; 0=Non-current smoker)

death 196 0.24 0.42 0 0 0 0 1 ttodth 196 2370 702 86 2074 2699 2826 2941 age 196 72.9 5.16 65 69 72 77 88 smoker 196 0.11 0.32 0 0 0 0 1	variable	N +	mean	sd	min	p25	p50	p75	max
	ttodth	196	2370	702	86	2074	2699	2826	2941
	age	196	72.9	5.16	65	69	72	77	88

The following three regression models (some output omitted) were run:

Model 1: Logistic regression adjusting for age and smoking status

. logistic death age smoker, robust

death	 Odds Ratio	Std. Err.			-	. Int.]
age smoker	1.203	0.04512 2.10962	4.95	0.000	1.118 1.2	

Appendix for Question 2 (cont.)

Model 2: Proportional Hazards Regression adjusting for age

```
. stset ttodth death
```

. stcox age, robust

_t		Robust Std. Err.	Z	P> z	[95% Conf. Interval]
age	1.034	0.0141	2.46	0.014	1.007 1.062

Model 3: Proportional Hazards Regression adjusting for age and smoking status

. stcox age smoker, robust

	 Haz. Ratio	Robust Std. Err.	z	P> z	[95% C	onf. Interval]
_	1.037	0.0144			1.009	_ , , , ,